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Ambient Air Pollutant Exposures and Hospitalization for Kawasaki Disease in Taiwan: A Case-Crossover Study (2000-2010)

Chau-Ren Jung^{1,2}, Wei-Ting Chen³, Yu-Ting Lin^{1,2}, and Bing-Fang Hwang¹

¹Department of Occupational Safety and Health, College of Public Health, China Medical University, Taichung, Taiwan; ²Department of Public Health, College of Public Health, China Medical University, Taichung, Taiwan; and ³Department of Atmospheric Science, National Taiwan University, Taipei, Taiwan

Correspondence with Bing-Fang Hwang, Department of Occupational Safety and Health, College of Public Health, China Medical University, No 91 Hsueh-Shih Rd, Taichung, Taiwan, 40402 R. O. C. [bfhwang@mail.cmu.edu.tw], Telephone: +886-4-22053366 ext 6208, Telefax: +886-4-22071861

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Abstract

Background: Kawasaki disease (KD) is an acute and multi-systemic vasculitis that occurs predominantly in infants and young children. Although the etiological agent of KD remains unclear, limited studies have reported that windborne environmental factors may trigger KD.

Objectives: We conducted a time-stratified case-crossover study to assess the associations between air pollutants and KD in Taiwan.

Methods: We identified children aged less than 5 years with a diagnosis of KD from the Longitudinal Health Insurance Database 2000 (LHID2000) between 2000 and 2010. We obtained data regarding carbon monoxide (CO), nitrogen dioxide (NO₂), ozone (O₃), particulate matter with aerodynamic diameter less than 10 μ m (PM₁₀), and sulfate dioxide (SO₂) from 70 monitoring stations and used inverse distance weighting to calculate average daily exposures for the residential postal code of each case. We performed conditional logistic regression to estimate associations between KD and each air pollutant according to interquartile range (IQR) increases and quartiles of exposure on the day of hospitalization versus 3-4 reference days during the same month for each case. Additionally, we estimated associations with single-day exposures lagged 1-2 days.

Results: We identified 695 KD hospital admissions during the study period. An IQR increase (28.73 ppb) of O₃ was positively associated with KD after adjusting for temperature, humidity, northward wind, and eastward wind (adjusted odds ratio = 1.21; 95% confidence interval 1.01, 1.44). There were no significant associations between KD and CO, NO₂, PM₁₀ or SO₂. The association with O₃ was limited to exposure on the day of hospitalization, and to exposure during the summer months (June-August).

Conclusions: Our results provide new evidence that exposure to O₃ may increase the risk of KD in children. However, further investigation is needed to confirm the association and identify a potential biological mechanism.

Introduction

Kawasaki disease (KD) is an acute and multi-systemic vasculitis that predominantly occurs in infants and young children aged less than 5 years. In approximately 15–25% of untreated cases, KD may affect the coronary arteries and lead to myocardial infarction (MI), sudden death or ischemic heart disease in childhood or early adulthood (Burns and Glodé 2004; Newburger et al. 2004). The clinical features of the disease are fever for more than five days, bilateral bulbar conjunctival injection without exudates, cervical lymphadenopathy, changes in extremities with both erythema and induration, changes in lips and oral mucosa including erythema, cracking, strawberry tongue, diffuse injection of oral and pharyngeal mucosa, as well as polymorphous rash (Newburger et al. 2004). Although several etiological theories have been proposed, including environmental toxin exposure, autoimmune pathogenesis and infectious diseases, the etiology of KD remains unknown (Rowley 2011).

The incidence of KD has a striking seasonal variation, with different patterns in several countries (Belay et al. 2006; Burns et al. 2005; Chang et al. 2004; Du et al. 2007; Huang et al. 2009; Lynch et al. 2003; Ng et al. 2005; Park et al. 2007), as well as geographical and temporal clustering (Kao et al. 2008). A study involving post-mortem examinations demonstrated that the causal agent of KD entered through the upper respiratory tract, leading to a systemic immune response in vascular tissue, pancreas and kidney (Rowley et al. 2000). Rodó and colleagues (2011) reported that an increase of KD cases in Japan, Hawaii and San Diego was associated with a large-scale shift in the Asia-North Pacific wind pattern, further suggesting that the etiological agent of KD is spread by wind. Further research conducted by this group suggested that the cause of KD in Japan may be an agent carried by tropospheric winds from the densely

cultivated region of northeastern China (Rodó et al. 2014). In addition, the authors concluded that the temporal pattern of KD cases is not consistent with an infectious etiology given evidence of a very short lag between exposure and the onset of symptoms (less than one day) (Rodó et al. 2014).

Ambient air contains a mixture of individual pollutants, such as free radicals, e.g., nitrogen dioxide (NO₂), and pollutants that have the ability to trigger free radical reactions, e.g., ozone (O₃) and particulate matter. Exposure to elevated air pollutants can stimulate oxidative stress, induce inflammation in the lung and vascular inflammation and cause subsequent responses that are particularly dangerous to susceptible individuals (Kelly 2003). To the best of our knowledge, only one study has assessed the associations between air pollutants and KD. In their analysis of KD in Japan, Rodó and colleagues (2014) also evaluated exposures to sulfur dioxide (SO₂), oxidants (O_x), nitrogen oxides (NO_x), carbon monoxide (CO), and non-methane hydrocarbon (measured at three stations in the Tokyo Metropolitan area) as potential causes of KD, but reported that results were negative (Rodó et al. 2014). Whether air pollutants contribute to KD is still unclear owing to a lack of consideration for possible confounding factors and limited monitoring data. We therefore conducted a case-crossover study to assess the associations between short-term air pollutants and KD after controlling for meteorological variables.

METHODS

Data source

The data used in the present study were sourced from the Longitudinal Health Insurance Database 2000 (LHID2000). LHID2000 is a representative subset of data that includes data from

1,000,000 individuals systematically and randomly selected from the year 2000 registry of beneficiaries of the Taiwan National Health Insurance Research Database (NHIRD) (NHRI 2016). NHIRD was established in March 1995 and includes detailed information, such as outpatient visit, hospital admission, prescriptions and diagnosis of disease based on the International Classification of Diseases, Ninth Revision, Clinical Modification (ICD-9-CM), from 99% of the entire population of 23 million enrollees in Taiwan. The National Health Insurance Administration (NHI) verifies the validity and quality of diagnosis by randomly sampling a constant ratio of claims from every hospital each year and through strict review by an independent group of medical experts (Lin et al. 2010). The accuracy of diagnosis of major diseases, such as acute coronary syndrome and ischemic stroke, has been validated (Cheng et al. 2011; Wu et al. 2010). Besides, KD is a catastrophic illness announced by NHI. Individuals who have a catastrophic illness certificate are free of charge for the premium and the co-payment for outpatient or inpatient care. The diagnosis of KD is assigned by a board-certified specialist, and needs to be reviewed and approved by NHI, therefore, the diagnosis of KD is accurate and reliable (Wei et al. 2014). National Health Research Institutes confirmed that there are no significant differences in the gender distribution, age distribution, number of newborns every year, and average insurance payroll amount between the beneficiaries of LHID2000 and NHIRD. Because the data were analyzed anonymously, the institute review board specifically waived the need for consent from each subject. This study has been approved by the Institute Review Board of China Medical University Hospital and it complies with the principles outlined in the Helsinki Declaration.

Outcome of interest

We identified children aged less than 5 years old who had a diagnosis of KD (ICD-9-CM code 446.1) from January 1, 2000 to December 31, 2010. To ensure the diagnostic validity of KD in the present study, only cases of KD where the ICD-9-CM code was assigned by a pediatrician were selected as our outcome of interest (Kuo et al. 2014). The index date of the KD cases was set as the date of hospital admission. We extracted 1,453 hospital admissions with a diagnosis of KD (ICD-9-CM code 446.1) from the LHID2000. We excluded patients more than 5 years old ($n = 756$), and patients with missing postcodes ($n = 2$).

For depicting the demographic characteristics of KD cases, municipal-level household income was accessed from annual average incomes data assessed by Taiwan Directorate-General of Budget, Accounting and Statistics. We divided Taiwan main island into four regions according to administrative divisions: northern, central, southern, and eastern region (Figure S1).

Exposure assessment

We obtained hourly data of CO, NO₂, SO₂, particulate matter with aerodynamic diameter less than 10 μ m (PM₁₀) and O₃ from 70 monitoring stations constructed by Taiwan Environmental Protection Administration (TEPA) on Taiwan's main island that provided measurements continuously from 2000 to 2010 (TEPA 2016). Methods used for measuring these pollutants were non-dispersive infrared absorption for CO, chemiluminescence for NO₂, ultraviolet absorption for O₃, beta-gauge for PM₁₀ and ultraviolet fluorescence for SO₂. These data were subjected to rigorous quality assurance and control procedure through independent projects.

TEPA authorized an independent private sector to perform annual performance audit and regular performance check for monitoring instruments (TEPA 2010).

For each monitoring stations, the daily average of air pollutants (24-hour for CO, NO₂, SO₂ and PM₁₀; 8-hour (10:00-18:00) for O₃) were calculated for subsequent analyses; at least 75% of the 1-hour values had to be available for the days included in this calculation. In addition, these 70 monitoring stations were classified into five types: 54 general stations, 4 industrial stations, 6 traffic stations, 2 national park stations, and 4 background stations. When daily average data of air pollutants for a monitoring station was not available, we used the average value from the other monitoring stations of the same type within the same day to fulfill missing values.

The locations of the monitoring stations and air pollution sources were identified and managed by a geographic information system (ArcGIS version 10; ESRI, Redlands, CA, USA). The monitoring data were interpolated to pollutant surfaces using the inverse distance weighting (IDW) method. For the IDW approach, we used suitable spatial resolution (100 meters) (Stroh et al. 2007) and the inverse squared distance (1/squared distance) weighted average of the three nearest monitors within 50 km of each grid cell (the study area was divided into 100 meters × 100 meters grid cells) to compute the daily average concentration for each air pollutant. Then, the daily air pollution data were assigned to individuals according to the residential postcode on the day of hospitalization. The average spatial resolution of postcodes in Taiwan was 17 ± 8.56 km² in urban areas, but was larger in rural areas (154 ± 104.39 km²) with low population density. For cross-validation of models, we randomly selected data from 63 monitoring stations (90% of 70 monitoring stations) to estimate air pollution using the IDW model and then retained seven stations (10%) for evaluation.

Meteorological variables

We used meteorological values from the European Center for Medium-Range Weather Forecasts (ECMWF)-Interim dataset (ECMWF 2015), a global atmospheric reanalysis from 1979 to current produced by the ECMWF (Dee et al. 2011). We extracted the following values from 2000 to 2010: temperature at 2 meters above the ground, dew point temperature at 2 meters above the ground, eastward wind at 10 meters above the ground (zonal wind or U wind), and northward wind at 10 meters above the ground (meridional wind or V wind), with a spatial resolution of 0.125° longitude \times 0.125° latitude in a nested grid covering Taiwan. We used values recorded at Coordinated Universal Time (UTC) 00:00, 06:00, 12:00 and 18:00 every day to calculate the daily average of these variables. The value of temperature and dew point temperature at 2 meters above the ground from the same 4 time points were used to calculate average relative humidity for each day. Owing to the moist environment in Taiwan (relative humidity $>50\%$), we used the following simple formula to compute relative humidity:

$$RH \approx 100 - 5(t - t_d)$$

where RH is the relative humidity in percent, and t_d and t are the dew point temperature at 2 meters above the ground and temperature at 2 meters above the ground in degrees Celsius, respectively (Lawrence 2005). Wind in the atmosphere can be represented mathematically by a vector to declare speed and direction and be decomposed into two components: eastward wind, whose positive value means wind flows from west to east; and northward wind, whose positive value means wind flows from south to north. The average daily values of meteorological variables were also assigned to individuals according to the postcodes.

Study design

We used a time-stratified case-crossover design to investigate the associations between exposure to air pollutants and the risk of KD. The case-crossover design is the most widely used design for investigating acute health effects of air pollution (Carracedo-Martínez et al. 2010). In this design, each individual serves as his or her own control by using the exposure on the days before or after the case day. Owing to the consistency of the subject across both days, this design can control for the influence of confounding factors such as gender, smoking history, occupational history and genetics (Maclure 1991). The time-stratified approach was proposed to avoid bias from season and day of the week by restricting the reference days to the same day of the week within the same month and year as the case day (the day of hospitalization for KD) (Janes 2005). For example, if a KD case was admitted to the hospital on the first Monday of January 2006 (January 2, 2006), all other Mondays within January 2006 were assigned as the reference days for that case (January 9, 16, 23, and 30, 2006). This approach resulted in three or four reference days for each case.

Statistical analyses

We used Spearman's correlation to examine the relationships between air pollutants and meteorological variables. Conditional logistic regression was used to estimate the associations between air pollutant and KD, with results reported as odds ratios (ORs) per interquartile range (IQR) with their 95% confidence intervals (CI) (PROC PHREG, release 9.4; SAS Institute, Cary, NC). We first fitted single pollutant models to estimate the association of exposure to an individual air pollutant with KD. If there was a statistically significant association (p -value < 0.05) between an individual pollutant and KD, we then examined the robustness of the

association between that specific pollutant and KD after controlling for other pollutants. If the correlation coefficient between two pollutants was greater than 0.5, we did not include it in the same model to avoid collinearity. We also compared the association of four exposure categories based on the quartiles of the distribution of an individual pollutant (high: >75th percentile; medium: 75th to 50th percentile; low: 50th to 25th percentile; and reference: <25th percentile) with the risk of KD. The chi-square test for linear trend in binomial proportions was applied to test whether there was an increasing trend in the proportion of KD for exposure categories. In addition to estimating associations with exposures on the day of hospitalization (lag 0), we estimated associations with single-day lagged exposures on the previous day (lag 1) and two days before hospitalization (lag 2). To assess any seasonal pattern, we conducted analyses for single pollutants stratified by season (spring: March–May; summer: June–August; autumn: September–November; winter: December–February). Daily mean values of temperature at 2 meters above the ground, relative humidity, eastward wind at the 10 meters above the ground, and northward wind at the 10 meters above the ground were included in all models as simple continuous covariates to adjust for potential confounding. Additionally, we conducted sensitivity analyses to compare ORs with and without adjustment for eastward and northward wind.

RESULTS

A total of 695 KD hospital admissions were identified from our study population from January 1 2000 to December 31 2010. The mean age at admission was 2.54 years, ranging from 0.11 to 4.99 years and most cases were male (56.83%) (Table 1). Approximately one-third of the children hospitalized for KD lived in municipalities with a high household income (>\$1,158,981 NTD; 34.10%). Most KD admissions were from the southern, northern, and western regions of

Taiwan (37.55%, 36.69%, and 23.74%, respectively); only a few were from the eastern region of Taiwan (2.01%) (Table 1).

Air pollution and meteorological variables

The IDW model performances were high for O₃ and PM₁₀ ($R^2 = 0.62$ for O₃; $R^2 = 0.78$ for PM₁₀) and low for CO, NO₂, and SO₂ ($R^2 = 0.20$ for CO; $R^2 = 0.32$ for NO₂; $R^2 = 0.35$ for SO₂) (Table S1). The distributions of daily average concentrations of air pollutants and meteorological variables by using time-stratified case-crossover design in Taiwan during the study period (2000–2010) are presented in Table 2. Annual average values of daily mean CO and NO₂ concentrations decreased during 2000–2010 (Table S2 and Figure S2). Annual average values of daily mean O₃ concentration increased steeply from 2000-2004 (from 40.14 ppb to 47.45 ppb), and then reached a plateau from 2005-2010 (fluctuated between 43.57 ppb to 48.07 ppb) (Table S2).

Table 3 shows the Spearman's correlation coefficients between daily air pollutant concentrations and meteorological variables. The daily average concentration of CO was highly and positively correlated with NO₂ ($r = 0.77$) and moderately correlated with PM₁₀ ($r = 0.48$). The concentration of PM₁₀ was highly and positively correlated with NO₂ ($r = 0.63$), and moderately correlated with O₃ ($r = 0.57$) and SO₂ ($r = 0.54$). The northward wind was moderately correlated with temperature ($r = 0.52$) and eastward wind ($r = 0.45$). In general, the daily air pollutant concentrations were negatively correlated with humidity (r ranged from -0.01 to -0.51) (Table 3).

The correlations between municipal-level household income and air pollution were weak in Taiwan ($r = -0.13$ for O_3 ; $r = 0.28$ for CO ; $r = 0.13$ for NO_2 ; $r = -0.09$ for PM_{10} ; $r = 0.01$ for SO_2)

Associations between air pollution and KD

In the single pollutant model, an IQR increase (per 28.73 ppb change) of O_3 concentration was positively associated with KD (adjusted OR = 1.21; 95% CI 1.01, 1.44) (Table 4). The second (27.18-41.13 ppb) and fourth (55.91-106.52 ppb) quartiles of O_3 exposure were significantly associated with KD when compared with the lowest quartile of exposure, but the OR for the third quartile was not significant, and we did not observe a clear linear trend when O_3 was modeled as a categorical variable (Table 4). The p -value for trend was significant for NO_2 modeled as a categorical variable, but the association was positive (but non-significant) for the highest quartile only (Table 4). There were no significant associations between KD and CO , PM_{10} , or SO_2 (Table 4). The association between KD and an IQR increase in O_3 remained stable and significant after adjusting for the other pollutants in two- and three-pollutant models (adjusted OR varying between 1.19 and 1.21) (Figure 1). However, we did not evaluate two- and three-pollutant models for associations with pollutants other than O_3 , because the other pollutants were not significantly associated with KD in single-pollutant models.

Figure 2 presents the lag pattern of air pollutants on KD. ORs for an IQR increase in CO and NO_2 were slightly higher for exposure on the previous day (lag 1) than on the same day (lag 0), while ORs for the other pollutants were positive for exposure on lag0 (the current day), close to the null or negative (<1.0) for exposure on the previous day (lag 1), and negative for exposure two days prior (lag 2) (Figure 2). The only significant positive association with KD was for O_3

on the same day. KD showed a significant negative association with IQR increases in O₃ and SO₂ at lag day 2 (OR = 0.85; 95% CI: 0.72, 1.00 and OR = 0.85; 95% CI: 0.72, 1.00, respectively).

Figure 3 shows the associations between air pollutants and KD stratified by season. For each pollutant, we used the same IQR value for all seasons. The association between O₃ and KD was significant in summer (adjusted OR = 1.50, 95% CI 1.05, 2.13 per 28.73 ppb), while the associations of O₃ for all other seasons were close to the null (Figure 3). In addition, the association between SO₂ and KD was significant in spring (adjusted OR = 1.11, 95% CI 1.02, 1.21 per 3.47 ppb change).

We also evaluated whether the eastward and northward wind at 10 meters above the ground might influence associations between air pollutants and KD. The ORs from model with and without adjusted for eastward and northward wind did not change more than 10 percent (Table S3).

DISCUSSION

In this study, we used a time-stratified case-crossover study to investigate the associations between air pollutants and KD. We estimated that an IQR increase (28.73 ppb) in O₃ concentration was associated with a 21% increase in the odds of hospitalization for KD (95% CI: 1, 44%). Thanks at least in part to more stringent air quality regulations and emission controls, there has been a long-term decrease in the traffic-related air pollutants CO and NO₂ (10.6 ppb and 0.21 ppb decrease per year) in Taiwan (Chen et al. 2015). Contrary to these gaseous pollutants, the concentration of O₃ has gradually increased (0.83 ppb increase per year) from

1994-2010 in Taiwan (Chen et al. 2015), which might represent a great environmental health problem. The increasing concentration of O₃ in western Taiwan reflects the economic growth in East Asia over the last few decades (Chen et al. 2015).

In the present study, we did not find a monotonic trend for the association between quartiles of O₃ with KD, though the association was strongest for the highest versus lowest quartile, positive for the second and third quartiles (though the OR for the third quartile was slightly lower than that for the second, and was not significant), and significant for an IQR increase in O₃ modeled as a continuous variable. After adjusting for temperature, humidity, northward wind and eastward wind, the significant negative associations of KD with O₃ and SO₂ exposures two days before hospitalization (lag 2, Figure 2) suggests a pattern consistent with short-term displacement (or “harvesting effect”) (Zanobetti and Schwartz 2008), whereby the children who are vulnerable to KD as a consequence of exposure develop symptoms on the day of exposure, leaving no vulnerable children left to develop KD 2 days later. However, this pattern needs to be confirmed in other study populations.

Although the mean O₃ concentration over Taiwan is lower during the summer than at other times of year due to the elevated mixing layer height, unstable atmospheric condition, and the formation of cloud or showers in afternoon during summer (Cheng et al. 2015; Chou et al. 2006; Tsai et al. 2008), the association between O₃ and KD appeared to be specific to exposures during the summer months (June-August). The observed results could be explained by the fact that children spend more time outside in summer and therefore could increase personal exposure level and there may be other explanations as well. However, additional researches will be required to confirm the seasonal pattern and investigate potential mechanisms.

Few studies have assessed the association between air pollution and KD. Rodó and colleagues postulated a fungus genus, *Candida*, as a possible factor that may be carried by wind from the cultivated region of northeastern China to Japan (Rodó et al. 2014). We did not have the data on the presence or amount of *Candida* in ambient air, but adjusting for eastward and northward wind as an indirect mean of controlling for potential confounding by fungus exposures in our analyses. Adjusting for both eastward and northward wind had little influence on associations (Table S3), which indicates that these variables were not strong confounders in our study. However, it is not possible for us to rule out potential confounding or effect modification by *Candida* exposure.

KD is an acute vasculitis that predominantly occurs in young children and particularly affects the coronary arteries (Brown et al. 2001; Burns and Glodé 2004; Newburger et al. 2004). Rowley and colleagues (2000) found a striking inflammation in the trachea of patients with KD. The progression of arterial lesions in KD involves perivasculitis of small arteries over days 0 to 9; panvasculitis of medium-sized, muscular arteries with aneurysm formation and thrombosis from days 12 to 25; coronary and other medium-sized artery myointimal proliferation from days 28 to 31; and narrowing of the arteries may occur after 40 days (Burns et al. 2000). Exposure to O₃ may enhance the release of inflammatory mediators from the airway epithelial cells and exacerbate the risk of adverse health effects in susceptible individuals (Bayram et al. 2001). Even in a developing primate infant, episodic exposure to O₃ can result in the loss of a number of terminal bronchioles, a reduction of distal airway size and alterations in smooth muscle bundle orientation (Fanucchi et al. 2006). Acute and chronic lung inflammation due to pollutants can result in systemic inflammation that triggers a cascade of events outside the lung and then causes

cardiovascular disease (Srebot et al. 2009; Tamagawa and van Eeden 2006). Chuang and colleagues (2009) used three different animal models (wild-type C57Bl/6 mice, apoE^{-/-} mouse, and infant macaque monkey) and found that inhalation of O₃ increased vascular dysfunction, oxidative stress and mitochondrial damage and, in apoE^{-/-} mouse, accelerated atherogenesis.

Several epidemiological studies have also reported associations between cardiovascular disease and O₃. Breton and colleagues (2012) conducted a cross-sectional study with a cohort of 768 American college students showed that exposure to a 9.3-ppb increase in O₃ during elementary school years (ages 6–12) was associated with a 10.1µm (95% CI: 1.8-18.5) higher carotid artery intima-media thickness in young adults. A case-crossover study involving 635 acute MI cases in France suggested that short-term exposure to a 5-ppb increase in O₃ on two consecutive days was related to a 5% increased risk of acute MI in French middle-aged adults (5%; 95% CI: 1-8% for the current day (lag 0), and 5%; 95% CI: 1-9% for the previous day (lag 1), respectively) (Ruidavets et al. 2005). A time series study conducted in Denver, one of the most polluted cities in the US, reported that exposure to an IQR change (9.7 ppb) in O₃ was positively associated with an increase in the risk of hospitalization for coronary atherosclerosis and pulmonary heart disease (12.3%, 95% CI: 4.0-21.4%, and 21.4%, 95% CI: 4.0-41.8%, respectively) (Koken et al. 2003). In a large meta-analysis involving 144 estimates from 39 time-series studies, Bell and colleagues (2005) found that a 10-ppb increase in O₃ was associated with a 2.45% (95%CI: 0–4.10%) increase in cardiovascular mortality during the warm season. By investigating 23 European cities with mean number of cardiovascular disease deaths per day ranged from 2 to 143 involved in Air Pollution and Health: a European Approach project (APHEA2), Gryparis and colleagues (2004) found a significant association between a 10 µg/m³

increases in 1-hr O₃ and cardiovascular mortality in the warm season (0.45%, 95% CI: 0.22–0.69%).

Our study had some potential limitations. First, air pollutant data were accessed from ambient air pollution monitoring stations rather than being measured at personal exposure levels, and could not represent indoor concentration of air pollutants. We used the IDW method in this study to estimate air pollution concentration based on postcode rather than personal address, which may have increased the possibility of random exposure misclassification. The IDW method is a more appropriate approach for secondary pollutants that vary on a larger geographical scale and are more homogeneously distributed (e.g., O₃) than for those that are more localized and dependent on combustible sources (e.g., CO, NO₂) (Vrijheid et al. 2011). In this study, we consistently found that the IDW model performed better for O₃ and PM₁₀ than for other pollutants (Table S1). The differences in associations between KD and the different pollutants is partly a function of the accuracy of exposure estimation. Second, possible confounders in studying the association of air pollution with KD include genetic factors, infectious disease (e.g., staphylococcus infection, superantigens), fungal toxins and socioeconomic factors, among others. Personal invariant factors such as genetic factors were controlled for using the time-stratified case-crossover design. Although possible confounding by socioeconomic factors could not be completely ruled out in our study, we did not observe a strong correlation between air pollution and municipal-level household income in Taiwan. The universal insurance program (NHIRD) in Taiwan was launched in 1995 and currently covers almost all residents (99%) (NHI 2012), therefore children's access to health services may be less influenced by socioeconomic factors in Taiwan than in some other countries. Additionally,

Chang and colleagues (2013) conducted a case-control study included 115 KD cases and 1,150 controls in Taiwan from 1997-2010 to assess the association between KD and urbanization. They used population density, percentage of people with college-level education or higher, percentage of elder aged large than 65 years, percentage of agricultural workers in the population, and number of physicians per 100,000 individuals in the population to classify urbanization and indicated that the association between urbanization and KD was not statistically significant. Unfortunately, we could not exclude confounding by infectious disease and fungal toxins due to a lack of data. Third, the index date (date of hospital admission) was the only available information on the timing of the KD event in our data source. The date of hospital admission does not represent the true timing of disease onset, which may have led to exposure misclassification. Although residents in Taiwan have variety of choice for seeking health care without financial barriers, easy access to the physicians and specialists, and short waiting times for health care services (Cheng 2015), the bias due to time mismatch between hospital admission and true timing of disease onset in our study was inevitable.

CONCLUSION

Our results provide new evidence that children exposed to O₃ may have an increased risk of hospitalization for KD. However, further investigation is needed to confirm our findings and evaluate potential biological mechanisms.

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Table 1. Demographic data of Kawasaki disease (KD) admissions (n=695, age 2.54 ± 1.21) from January 1 2000 to December 31 2010 in Taiwan.

Variables	Number of events (%)
Sex	
Female	300 (43.17)
Male	395 (56.83)
Household income	
<\$897,000 NTD	198 (28.49)
\$897,000–1,095,192 NTD	136 (19.57)
\$1,095,192–1,158,981 NTD	124 (17.84)
>\$1,158,981 NTD	237 (34.10)
Region	
Northern	255 (36.69)
Central	165 (23.74)
Southern	261 (37.55)
Eastern	14 (2.01)

Note: Municipal-level household income of KD admissions were accessed from annual average incomes data assessed by Taiwan Directorate-General of Budget, Accounting and Statistics. The regions of Taiwan were divided according to administrative divisions.

Abbreviations: NTD, New Taiwan Dollar.

Table 2. Distribution of daily data of air pollutants and meteorological variables by using time-stratified case-crossover design in Taiwan, 2000–2010.

Daily Environmental Variable	Mean ± SD	Minimum	Q1	Q2	Q3	Maximum	IQR
Air pollutants							
CO (ppm)	0.75 ± 0.35	0.15	0.54	0.69	0.88	9.27	0.34
NO ₂ (ppb)	23.20 ± 9.41	2.64	16.02	22.33	29.36	72.02	13.34
O ₃ (ppb)	42.75 ± 19.36	1.17	27.18	41.13	55.91	106.52	28.73
PM ₁₀ (µg/m ³)	57.39 ± 29.88	10.75	34.54	50.62	75.14	235.43	40.60
SO ₂ (ppb)	4.43 ± 2.80	0.05	2.40	3.92	5.87	25.68	3.47
Meteorological variable							
Temperature (°C)	21.86 ± 5.03	5.54	18.14	22.89	25.97	30.69	7.83
Humidity (%)	81.85 ± 8.67	51.15	75.86	82.38	88.36	98.98	12.50
U10 (m/s)	-1.10 ± 2.02	-10.76	-2.46	-1.19	0.19	7.95	2.65
V10 (m/s)	-0.72 ± 2.19	-9.88	-2.26	-0.75	0.86	8.60	3.12

Abbreviations: CO, carbon monoxide; IQR: interquartile range; NO₂, nitrogen dioxide; O₃, ozone; PM₁₀, particulate matter with aerodynamic diameter equal to or less than 10 µm; Q1, 25th percentile; Q2, 50th percentile; Q3, 75th percentile; SD, standard deviation; SO₂, sulfur dioxide; U10, eastward wind at 10 meters above the ground; V10, northward wind at 10 meters above the ground.

Table 3. Spearman's correlation between daily average air pollutants and meteorological variables in Taiwan, 2000–2010.

	CO	NO ₂	O ₃	PM ₁₀	SO ₂	Temperature	Humidity	U10
CO	1.00							
NO ₂	0.77**	1.00						
O ₃	0.14**	0.23**	1.00					
PM ₁₀	0.48**	0.63**	0.57**	1.00				
SO ₂	0.40**	0.55**	0.27**	0.54**	1.00			
Temperature	-0.23**	-0.43**	0.05*	-0.30**	0.02	1.00		
Humidity	-0.01	-0.17**	-0.51**	-0.43**	-0.29**	-0.05*	1.00	
U10	-0.03	-0.11**	-0.10**	-0.14**	-0.01	0.39**	0.11**	1.00
V10	-0.07**	-0.22**	-0.09**	-0.21**	-0.00	0.52**	-0.01	0.45**

Abbreviations: CO, carbon monoxide; NO₂, nitrogen dioxide; O₃, ozone; PM₁₀, particulate matter with aerodynamic diameter equal to or less than 10 μm; SO₂, sulfur dioxide; U10, eastward wind at 10 meters above the ground; V10, 10-m northward wind at 10 meters above the ground.

* $p < 0.05$; ** $p < 0.001$

Table 4. Associations between Kawasaki disease and air pollution from single pollutant models in Taiwan, 2000–2010.

Air pollutants	OR (95% CI)	Adjusted OR ^b	<i>p</i> for trend ^c
O₃			
per 28.73 ppb change	1.17 (1.01, 1.35)	1.21 (1.01, 1.44)	
1st Quartile ^a , 1.17–27.18 ppb	1.00	1.00	0.11
2nd Quartile, 27.18–41.13 ppb	1.27 (0.99, 1.63)	1.30 (1.00, 1.68)	
3rd Quartile, 41.13–55.91 ppb	1.23 (0.95, 1.59)	1.28 (0.97, 1.70)	
4th Quartile, 55.91–106.52 ppb	1.33 (1.00, 1.75)	1.40 (1.01, 1.94)	
NO₂			
per 13.34 ppb change	1.05 (0.87, 1.26)	1.08 (0.89, 1.30)	
1st Quartile ^a , 2.64–16.02 ppb	1.00	1.00	0.02
2nd Quartile, 16.02–22.33 ppb	0.99 (0.75, 1.30)	0.98 (0.74, 1.29)	
3rd Quartile, 22.33–29.36 ppb	1.00 (0.72, 1.39)	0.99 (0.71, 1.38)	
4th Quartile, 29.36–72.02 ppb	1.24 (0.87, 1.76)	1.28 (0.90, 1.83)	
CO			
per 0.34 ppm change	0.99 (0.89, 1.11)	1.01 (0.91, 1.13)	
1st Quartile ^a , 0.15–0.54 ppm	1.00	1.00	0.06
2nd Quartile, 0.54–0.69 ppm	0.91 (0.70, 1.19)	0.92 (0.70, 1.19)	
3rd Quartile, 0.69–0.88 ppm	1.03 (0.77, 1.37)	1.04 (0.78, 1.40)	
4th Quartile, 0.88–9.27 ppm	1.12 (0.81, 1.54)	1.20 (0.86, 1.66)	
PM₁₀			
per 40.60 µg/m ³ change	1.10 (0.94, 1.28)	1.10 (0.94, 1.30)	
1st Quartile ^a , 10.75–34.54 µg/m ³	1.00	1.00	0.22
2nd Quartile, 34.54–50.62 µg/m ³	1.12 (0.87, 1.45)	1.11 (0.86, 1.45)	
3rd Quartile, 50.62–75.14 µg/m ³	1.01 (0.77, 1.33)	1.01 (0.75, 1.35)	
4th Quartile, 75.14–235.43 µg/m ³	1.17 (0.85, 1.60)	1.16 (0.83, 1.63)	
SO₂			
per 3.47 ppb change	1.06 (0.92, 1.22)	1.06 (0.92, 1.23)	
1st Quartile ^a , 0.05–2.40 ppb	1.00	1.00	0.69
2nd Quartile, 2.40–3.92 ppb	1.19 (0.91, 1.54)	1.18 (0.82, 1.55)	
3rd Quartile, 3.92–5.87 ppb	1.09 (0.82, 1.45)	1.10 (0.82, 1.48)	
4th Quartile, 5.87–25.68 ppb	1.13 (0.83, 1.55)	1.14 (0.82, 1.59)	

Abbreviations: CO, carbon monoxide; NO₂, nitrogen dioxide; O₃, ozone; OR, odds ratio; PM₁₀, particulate matter with aerodynamic diameter equal to or less than 10 µm; SO₂, sulfur dioxide.

^aReference group.

^bAdjusted ORs were adjusted for temperature, humidity, northward wind, and eastward wind.

^cThe *p* for trend was computed by using the chi-square test for linear trend in binomial proportions.

Figure 1. Associations between Kawasaki disease and ozone from two-and three-pollutant models in Taiwan, 2000–2010. The associations are showed as odds ratio (OR) with 95% confidence interval (CI) for an interquartile range (28.73 ppb) increases in ozone. All models were adjusted for temperature, humidity, northward wind, and eastward wind. Model 1 included ozone and carbon monoxide. Model 2 included ozone and nitrogen dioxide. Model 3 included ozone and sulfur dioxide. Model 4 included ozone, sulfur dioxide, and carbon monoxide. Model 5 included ozone, sulfur dioxide, and nitrogen dioxide. CI, confidence interval.

Figure 2. Associations between Kawasaki disease and air pollution in Taiwan, 2000–2010 at different lag days. The associations are showed as odds ratio (OR) with 95% confidence interval (CI) for interquartile range (IQR) increases in exposure (see Table 2 for the IQRs for each pollutant). All models were adjusted for temperature, humidity, northward wind, and eastward wind. CO, carbon monoxide; NO₂, nitrogen dioxide; O₃, ozone; PM₁₀, particulate matter with aerodynamic diameter equal to or less than 10 µm; SO₂, sulfur dioxide.

Figure 3. Associations between air pollutants and Kawasaki disease stratified by season. The associations are showed as odds ratio (OR) with 95% confidence interval (CI) for interquartile range (IQR) increase in exposure (the IQRs are based on overall distributions, see Table 2 for the IQRs for each pollutant). Models were adjusted for temperature, humidity, eastward wind, and northward wind. Autumn, September–November; CO, carbon monoxide; NO₂, nitrogen dioxide; O₃, ozone; PM₁₀, particulate matter with aerodynamic diameter equal to or less than 10 µm; SO₂, sulfur dioxide; Spring, March–May; Summer, June–August; Winter, December–February.

Figure 1.

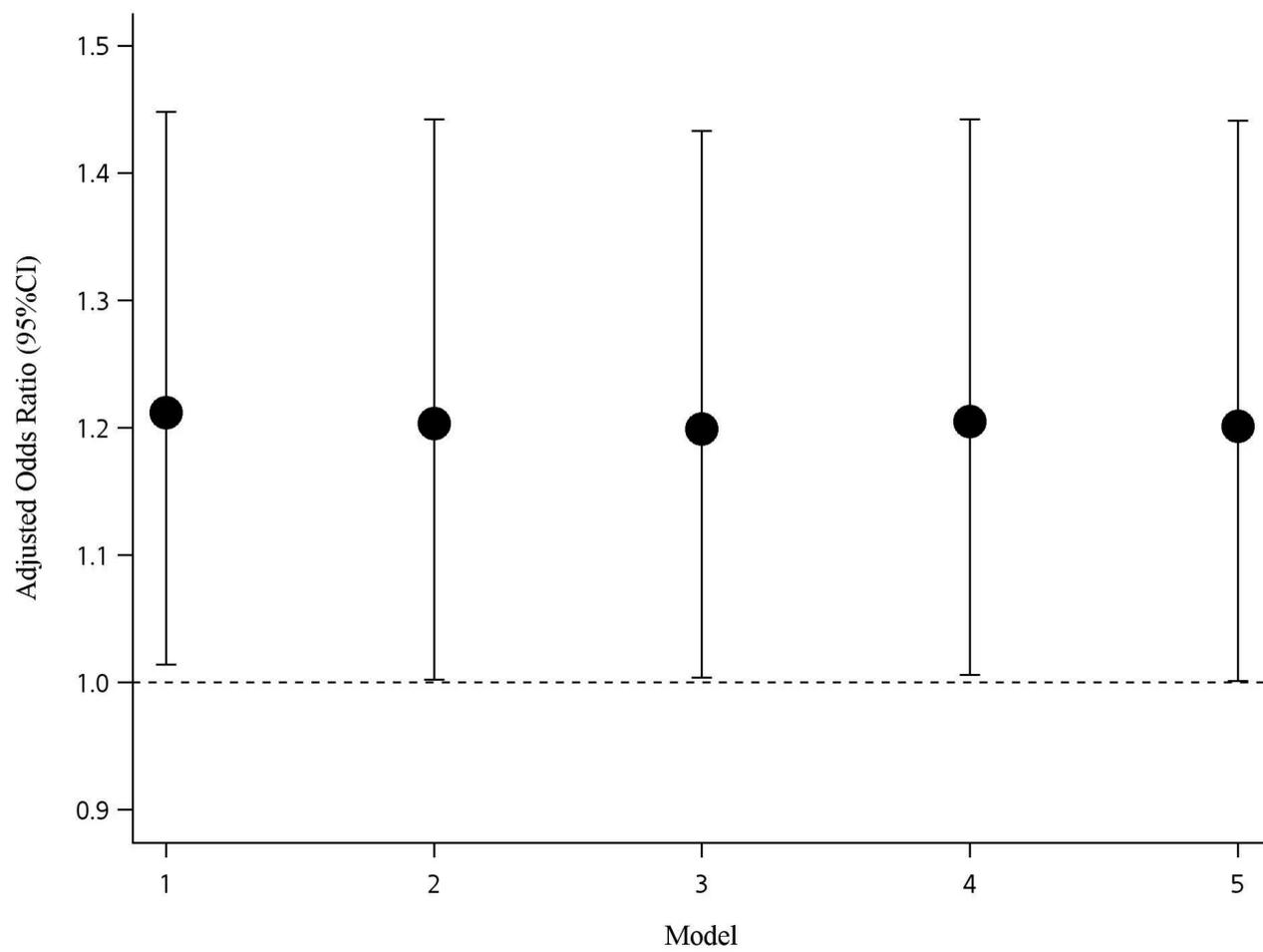


Figure 2.

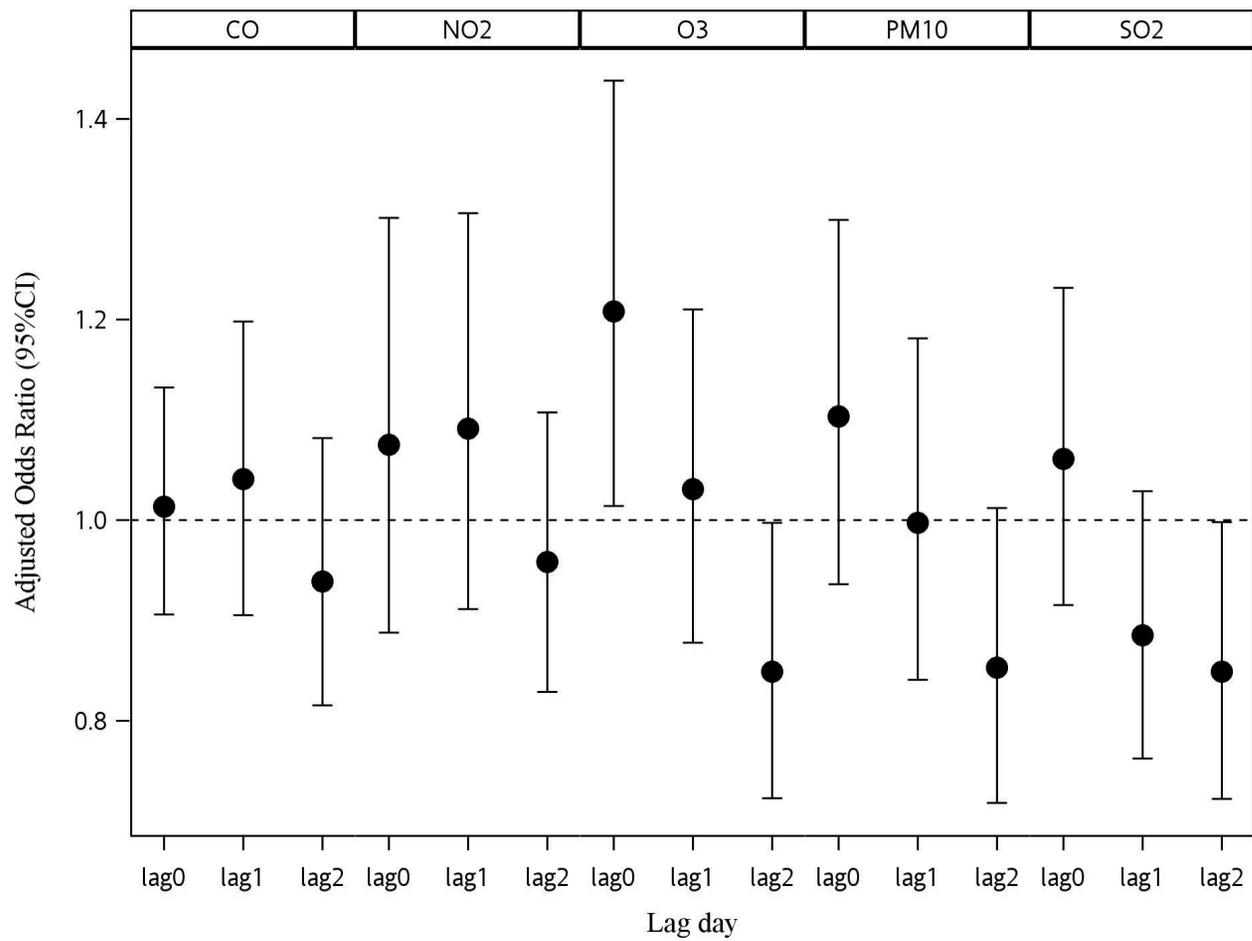


Figure 3.

